

$$x^2 - 290x = 289\sqrt{x}, \text{ compute } \overline{\sqrt{x} - \sqrt{x}}$$

$$x^2 - x = (x + \sqrt{x})(x - \sqrt{x})$$

$$x^2 - x = 289(\sqrt{x} + x)$$

$$(x + \sqrt{x})(x - \sqrt{x}) = 289(x + \sqrt{x})$$

$$x - \sqrt{x} = 289$$

$$\overline{\sqrt{x} - \sqrt{x}} = 17$$

$$X_1, X_2, \dots \text{ i.i.d. } E|X_i| < \infty \Rightarrow \lim_{n \rightarrow \infty} \bar{X}_n \stackrel{a.s.}{=} E X_1 \quad (\text{SLLN})$$

$\leftarrow \mu$

$$E X_i^2 < \infty, \text{ then } \lim_{n \rightarrow \infty} \sqrt{n}(\bar{X}_n - \mu) \stackrel{d}{=} N(0, \sigma^2) \quad (\text{CLT})$$

$\leftarrow \text{Var}(X_1)$

$$\text{Slutsky: } X_n \xrightarrow{d} X$$

$$Y_n \xrightarrow{d} \alpha = \text{constant} \neq 0 \Rightarrow \begin{cases} X_n \pm Y_n \xrightarrow{d} X \pm \alpha \\ \frac{X_n}{Y_n} \xrightarrow{d} \frac{X}{\alpha} \\ X_n Y_n \xrightarrow{d} X \alpha \end{cases}$$

$$\text{Ex. 1) } \sqrt{n}(\hat{\sigma}_n^2 - \sigma^2) \xrightarrow{d} N(0, 2)$$

2) approx $100(1-\alpha)\%$ CI for μ when σ^2 is unknown

$$\bar{X}_n \pm \frac{S_n}{\sqrt{n}} z_{\alpha/2}$$

$$\text{why? } \sqrt{n}(\bar{X}_n - \mu) \xrightarrow{d} N(0, \sigma^2)$$

$$\frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma} \xrightarrow{d} N(0, 1)$$

$$\text{SLLN } S_n^2 \xrightarrow{a.s.} \sigma^2 > 0 \quad \text{slutsky } \frac{\sqrt{n}(\bar{X}_n - \mu)}{S_n} \xrightarrow{d} N(0, 1)$$

Def: 1) $\hat{\theta}_n$ is consistent if $\lim_{n \rightarrow \infty} \hat{\theta}_n \stackrel{a.s.}{=} \theta$
(strongly)

2) $\hat{\theta}_n$ is asymptotically normal if $\exists a_n \rightarrow \infty$, s.t.
 $a_n(\hat{\theta}_n - \theta) \xrightarrow{d} N(0, 1)$

asym normal \Rightarrow consistent

Ex. (not asymp normal)

$$X \sim U(0, \theta)$$

$$\hat{\theta}_n = X_{(n)} < \theta$$

$n(\theta - X_{(n)}) \xrightarrow{d}$ exponential

$$\begin{aligned} P(n(\theta - X_{(n)}) > x) &= P(\theta - X_{(n)} > \frac{x}{n}) = P(X_{(n)} < \theta - \frac{x}{n}) \\ &= P\left(\frac{X_{(n)}}{\theta} < 1 - \frac{x}{n\theta}\right) = \left(1 - \frac{x}{n\theta}\right)^n \xrightarrow{n \rightarrow \infty} e^{-x/\theta} \\ &\quad \text{" } U_{(n)} \qquad \qquad \qquad \text{" } P(Y > x) \end{aligned}$$

δ -method: if we have $a_n(\hat{\theta}_n - \theta) \xrightarrow{d} N(0, 1)$

(delta) and $f = f(x)$ $f'(\theta) \neq 0$

then $a_n(f(\hat{\theta}_n) - f(\theta)) \xrightarrow{d} N(0, |f'(\theta)|^2)$

why?

Taylor

$$f(\hat{\theta}_n) = f(\theta) + f'(\tilde{\theta}_n)(\theta - \hat{\theta}_n)$$

$$\hat{\theta}_n - \theta \rightarrow 0$$

$$f(\hat{\theta}_n) - f(\theta) = \underbrace{f'(\tilde{\theta}_n)}_{\xrightarrow{n \rightarrow \infty} f'(\theta)} (\theta - \hat{\theta}_n)$$

$$\Leftrightarrow \tilde{\theta}_n \rightarrow \theta$$



$$f'(\theta) N(0, 1)$$

$$\xrightarrow{d} N(0, 1)$$